High-Dimensional Sparse Surrogate Construction via Bayesian Compressive Sensing

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OUTLINE

- Surrogates needed for complex models
- Polynomial Chaos (PC) surrogates do well with uncertain inputs
- Bayesian regression provide results with uncertainty certificate
- Compressive sensing ideas deal with high-dimensionality

Surrogate construction: scope and challenges

Construct surrogate for a complex model $f(\lambda)$ to enable

- Global sensitivity analysis
- Optimization
- Forward uncertainty propagation
- Input parameter calibration
- • •

- Computationally expensive model simulations, data sparsity
 - Need to build accurate surrogates with as few training runs as possible
- High-dimensional input space
 - Too many samples needed to cover the space
 - Too many terms in the polynomial expansion

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Build/presume PC for input parameter λ

$$\lambda(\mathbf{x}) = \sum_{k=0}^{K-1} \mathbf{a}_k \Psi_k(\mathbf{x})$$

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• E.g., gaussian with known moments μ_i , σ_i ,

$$\lambda_i = \mu_i + \sigma_i x_i$$

Build/presume PC for input parameter λ

$$\lambda(\mathbf{x}) = \sum_{k=0}^{K-1} \mathbf{a}_k \Psi_k(\mathbf{x})$$

• Input parameters are represented via their cumulative distribution function $F(\cdot)$, such that, with $x_i \sim \text{Uniform}[-1, 1]$

$$\lambda_i = F_{\lambda_i}^{-1} \left(\frac{x_i + 1}{2} \right), \quad \text{for } i = 1, 2, \dots, d.$$

Build/presume PC for input parameter λ

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• If input parameters are uniform in $[a_i, b_i]$, then

$$\lambda_i = \frac{a_i + b_i}{2} + \frac{b_i - a_i}{2} x_i$$

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• Forward function $f(\cdot)$, output u

$$u = f(\lambda(\mathbf{x}))$$

$$u = \sum_{k=0}^{K-1} c_k \Psi_k(\mathbf{x}) \equiv g(\mathbf{x})$$

- Global sensitivity information for free
 - Sobol indices, variance-based decomposition.

Alternative methods to obtain PC coefficients

$$u \simeq \sum_{k=0}^{K-1} c_k \Psi_k(\boldsymbol{x})$$

• <u>Projection</u> $c_k = \frac{\langle u(\mathbf{x})\Psi_k(\mathbf{x})\rangle}{\langle \Psi_k^2(\mathbf{x})\rangle}$ The integral $\langle u(\mathbf{x})\Psi_k(\mathbf{x})\rangle = \int u(\mathbf{x})\Psi_k(\mathbf{x})d\mathbf{x}$ can be estimated by

Monte-Carlo

$$\frac{1}{N}\sum_{i=1}^{N}u(\mathbf{x}_{i})\Psi_{k}(\mathbf{x}_{i})$$



many(!) random samples

Quadrature

$$\sum_{j=1}^{Q} u(\mathbf{x}_j) \Psi_k(\mathbf{x}_j) w_j$$



samples at quadrature

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 $many (!) \ random \ samples$

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samples at quadrature

Bayesian regression

$$P(c_k|u(\mathbf{x}_j)) \propto P(u(\mathbf{x}_j)|c_k)P(c_k)$$



any (number of) samples

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Bayesian regression

$$\underline{P(c|\mathcal{D})} \propto \underline{P(\mathcal{D}|c)} \underline{P(c)}$$
Posterior Likelihood Prior

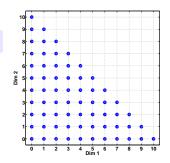


any (number of) samples

$$y = u(\mathbf{x}) \approx \sum_{k=0}^{K-1} c_k \Psi_k(\mathbf{x})$$

$$\Psi_k(x_1, x_2, ..., x_d) = \psi_{k_1}(x_1)\psi_{k_2}(x_2)\cdots\psi_{k_d}(x_d)$$

- Issues:
 - how to properly choose the basis set?

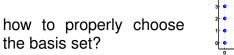


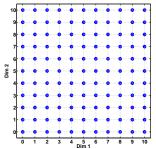
- need to work in underdetermined regime N < K: fewer data than bases (d.o.f.)
- Discover the underlying low-d structure in the model
 - get help from the machine learning community

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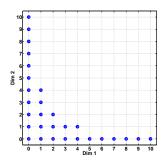


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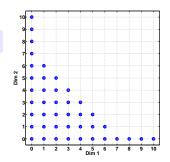


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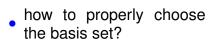


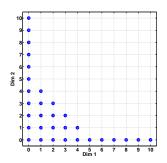
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In a different language....

- *N* training data points (x_n, u_n) and *K* basis terms $\Psi_k(\cdot)$
- Projection matrix $\mathbf{P}^{N \times K}$ with $\mathbf{P}_{nk} = \Psi_k(\mathbf{x}_n)$
- Find regression weights $c = (c_0, \dots, c_{K-1})$ so that

$$u \approx Pc$$
 or $u_n \approx \sum_k c_k \Psi_k(\mathbf{x}_n)$

- The number of polynomial basis terms grows fast; a p-th order, d-dimensional basis has a total of K = (p+d)!/(p!d!) terms.
- For limited data and large basis set (N < K) this is a sparse signal recovery problem ⇒ need some regularization/constraints.
- Least-squares $argmin_{c} \{||u-Pc||_{2}\}$
- The 'sparsest' $\operatorname{argmin}_{\boldsymbol{c}} \left\{ ||\boldsymbol{u} \boldsymbol{P} \boldsymbol{c}||_2 + \alpha ||\boldsymbol{c}||_0 \right\}$
- ullet Compressive sensing $\mathit{argmin}_{oldsymbol{c}}\left\{||oldsymbol{u}-oldsymbol{P}oldsymbol{c}||_{2}+lpha||oldsymbol{c}||_{1}
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- Compressive sensing $\mathit{argmin}_{\pmb{c}} \; \{ ||\pmb{u} \pmb{Pc}||_2 + \alpha ||\pmb{c}||_1 \}$ Bayesian Likelihood Prior

Weighted Bayesian Compressive Sensing

Dimensionality reduction by using hierarchical priors

$$p(c_k|\sigma_k^2) = \frac{1}{\sqrt{2\pi}\sigma_k} e^{-\frac{c_k^2}{2\sigma_k^2}} \qquad \qquad p(\sigma_k^2|\alpha_k) = \frac{\alpha_k}{2} e^{-\frac{\alpha_k\sigma_k^2}{2}}$$

Effectively, one obtains Laplace sparsity prior

$$p(\boldsymbol{c}|\boldsymbol{\alpha}) = \int \prod_{k=0}^{K-1} p(c_k|\sigma_k^2) p(\sigma_k^2|\alpha_k) d\sigma_k^2 = \prod_{k=0}^{K-1} \frac{\sqrt{\alpha_k}}{2} e^{-\sqrt{\alpha_k}|c_k|}$$

• Evidence maximization dictates values for $\sigma_k^2, \alpha_k, \sigma^2$ and allows exact Bayesian solution

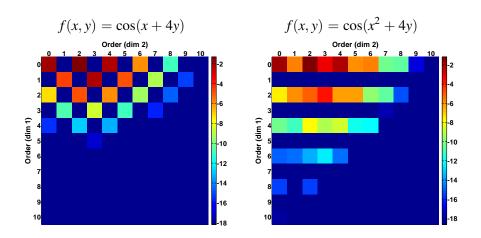
$$c \sim \mathcal{MVN}(oldsymbol{\mu}, oldsymbol{\Sigma})$$

with

$$\mu = \sigma^{-2} \Sigma P^T u$$
 $\Sigma = \sigma^2 (P^T P + \text{diag}(\sigma^2 / \sigma_k^2))^{-1}$

• KEY: Some $\sigma_k^2 \to 0$, hence the corresponding basis terms are dropped.

BCS removes unnecessary basis terms



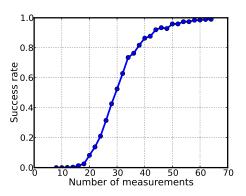
The square (i,j) represents the (log) spectral coefficient for the basis term $\psi_i(x)\psi_i(y)$.

Success rate grows with more data and 'sparser' model

Consider test function

$$f(\mathbf{x}) = \sum_{k=0}^{K-1} c_k \Psi_k(\mathbf{x})$$

where only S coefficients c_k are non-zero. Typical setting is

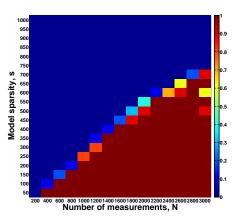


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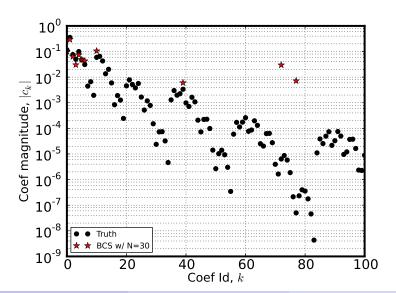
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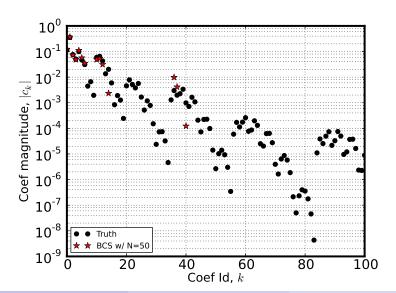
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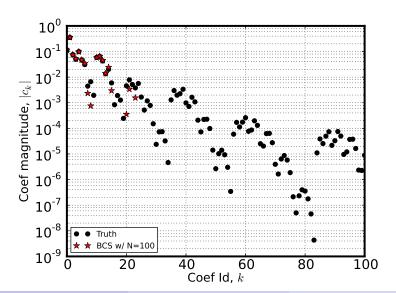
BCS recovers true PC coefficients with increased number of measurements



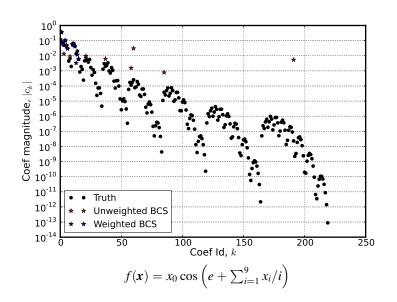
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BCS recovers true PC coefficients with increased number of measurements



WBCS recovers true coefficients better



Iteratively reweighting Compressive Sensing

[Candes et al., 2007]

Sparsest solution: $min||c||_0$ such that $u \approx Pc$

Compressive sensing: $min||c||_1$ such that $u \approx Pc$

Weighted compressive sensing: $min||Wc||_1$ such that $u \approx Pc$

Sparsest solution:

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Weighted compressive sensing:

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For sparse signals, $u = Pc^s$, with $||c_s||_0 = S < K$, ideal weights are

$$m{W} = diag\left(rac{1}{|c_k^s|}
ight)$$
 [i.e., $W_{kk} = +\infty$ if $c_k^s = 0$]

In practice, the true signal coefficients are not known, so...

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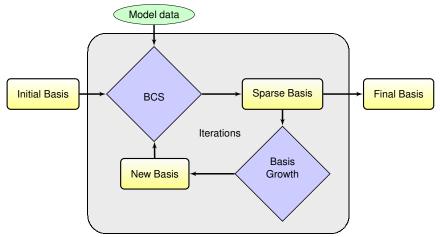
Iterative re-weighting

$$\mathbf{W}^{(i+1)} = diag\left(\frac{1}{|c_k^{(i)}| + \epsilon}\right)$$

 $[\epsilon \ll 1 \text{ for stability}]$

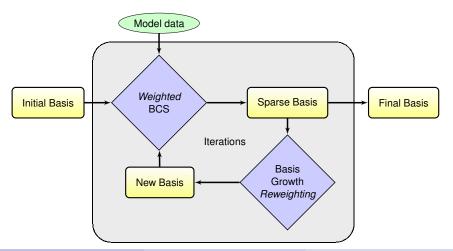
Iterative Bayesian Compressive Sensing (iBCS)

 Iterative BCS: We implement an iterative procedure that allows increasing the order for the relevant basis terms while maintaining the dimensionality reduction [Sargsyan et al. 2014], [Jakeman et al. 2015].



Iterative Bayesian Compressive Sensing (iBCS)

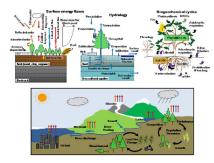
Combine basis growth and reweighting!



Basis set growth: simple anisotropic function

Basis set growth: ... added outlier term

Application of Interest: Community Land Model

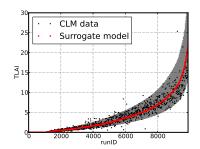


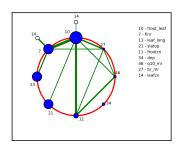
http://www.cesm.ucar.edu/models/clm/

- Nested computational grid hierarchy
- ullet A single-site, 1000-yr simulation takes ~ 10 hrs on 1 CPU
- Involves ∼ 50 input parameters; some dependent
- Non-smooth input-output relationship

Sparse PC surrogate for the Community Land Model

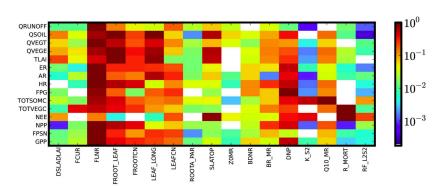
- Main effect sensitivities : rank input parameters
- Joint sensitivities : most influential input couplings
- About 200 polynomial basis terms in the 50-dimensional space
- Sparse PC will further be used for
 - sampling in a reduced space
 - parameter calibration against experimental data





Sparse PC surrogate for the Community Land Model

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Summary

- Surrogate models are necessary for complex models
 - Replace the full model for both forward and inverse UQ
- Uncertain inputs
 - Polynomial Chaos surrogates well-suited
- Limited training dataset
 - Bayesian methods handle limited information well
- Curse of dimensionality
 - The hope is that not too many dimensions matter
 - Compressive sensing (CS) ideas ported from machine learning
 - We implemented iteratively reweighting Bayesian CS algorithm that reduces dimensionality and increases order on-the-fly.

- Open issues
 - Computational design. What is the best sampling strategy?
 - Overfitting still present. Cross-validation techniques help.

Literature

- M. Tipping, "Sparse Bayesian learning and the relevance vector machine", J Machine Learning Research, 1, pp. 211-244, 2001.
- S. Ji, Y. Xue and L. Carin, "Bayesian compressive sensing", IEEE Trans. Signal Proc., 56:6, 2008.
- S. Babacan, R. Molina and A. Katsaggelos, "Bayesian compressive sensing using Laplace priors", IEEE Trans. Image Proc., 19:1, 2010.
- E. J. Candes, M. Wakin and S. Boyd. "Enhancing sparsity by reweighted ℓ₁ minimization", J. Fourier Anal. Appl., 14 877-905, 2007.
- A. Saltelli, "Making best use of model evaluations to compute sensitivity indices", Comp Phys Comm, 145, 2002.
- K. Sargsyan, C. Safta, H. Najm, B. Debusschere, D. Ricciuto and P. Thornton, "Dimensionality reduction for complex models via Bayesian compressive sensing", Int J for Uncertainty Quantification, 4(1), pp. 63-93,2014.
- J. Jakeman, M. Eldred and K. Sargsyan, "Enhancing ℓ_1 -minimization estimates of polynomial chaos expansions using basis selection", *J Comp Phys*, in press, 2015, see ArXiv.

Random variables represented by Polynomial Chaos

$$X \simeq \sum_{k=0}^{K-1} c_k \Psi_k(\boldsymbol{\eta})$$

• $\eta = (\eta_1, \dots, \eta_d)$ standard i.i.d. r.v. Ψ_k standard polynomials, orthogonal w.r.t. $\pi(\eta)$.

$$\Psi_k(\eta_1, \eta_2, \dots, \eta_d) = \psi_{k_1}(\eta_1)\psi_{k_2}(\eta_2)\cdots\psi_{k_d}(\eta_d)$$

- Typical truncation rule: total-order $p, k_1 + k_2 + \dots k_d \le p$. Number of terms is $K = \frac{(d+p)!}{d! p!}$.
- Essentially, a parameterization of a r.v. by deterministic spectral modes c_k .
- Most common standard Polynomial-Variable pairs: (continuous) Gauss-Hermite, <u>Legendre-Uniform</u>, (discrete) Poisson-Charlier.

Bayesian inference of PC surrogate

$$u \simeq \sum_{k=0}^{K-1} c_k \Psi_k(\mathbf{x}) \equiv g_{\mathbf{c}}(\mathbf{x})$$
 Posterior Likelihood Prior $P(\mathbf{c}|\mathcal{D}) \propto P(\mathcal{D}|\mathbf{c})$ Provided Prior $P(\mathbf{c}|\mathcal{D}) \propto P(\mathcal{D}|\mathbf{c})$ Provided Prior $P(\mathbf{c}|\mathcal{D}) \propto P(\mathcal{D}|\mathbf{c})$ Prior $P(\mathcal{D}|\mathbf{c}) \sim P(\mathcal{D}|\mathbf{c})$ Prior $P(\mathbf{c}|\mathcal{D}) \propto P(\mathcal{D}|\mathbf{c})$

• Data consists of training runs

$$\mathcal{D} \equiv \{(\boldsymbol{x}_i, u_i)\}_{i=1}^N$$

• <u>Likelihood</u> with a gaussian noise model with σ^2 fixed or inferred,

$$L(\boldsymbol{c}) = P(\mathcal{D}|\boldsymbol{c}) = \left(\frac{1}{\sigma\sqrt{2\pi}}\right)^N \prod_{i=1}^N \exp\left(-\frac{(u_i - g\boldsymbol{c}(\boldsymbol{x}))^2}{2\sigma^2}\right)$$

- Prior on c is chosen to be conjugate, uniform or gaussian.
- Posterior is a multivariate normal

$$oldsymbol{c} \in \mathcal{MVN}(oldsymbol{\mu},oldsymbol{\Sigma})$$

The (uncertain) surrogate is a gaussian process

$$\sum_{k=0}^{K-1} c_k \Psi_k(\pmb{x}) = \pmb{\Psi}(\pmb{x})^T \pmb{c} \quad \in \quad \mathcal{GP}(\pmb{\Psi}(\pmb{x})^T \pmb{\mu}, \pmb{\Psi}(\pmb{x}) \pmb{\Sigma} \pmb{\Psi}(\pmb{x}')^T)$$

Sensitivity information comes free with PC surrogate,

$$g(x_1,\ldots,x_d)=\sum_{k=0}^{K-1}c_k\Psi_k(\mathbf{x})$$

Main effect sensitivity indices

$$S_{i} = \frac{Var[\mathbb{E}(g(\mathbf{x}|x_{i}))]}{Var[g(\mathbf{x})]} = \frac{\sum_{k \in \mathbb{I}_{i}} c_{k}^{2} ||\Psi_{k}||^{2}}{\sum_{k > 0} c_{k}^{2} ||\Psi_{k}||^{2}}$$

 \mathbb{I}_i is the set of bases with only x_i involved

Sensitivity information comes free with PC surrogate,

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Joint sensitivity indices

$$S_{ij} = \frac{Var[\mathbb{E}(g(\mathbf{x}|x_i, x_j))]}{Var[g(\mathbf{x})]} - S_i - S_j = \frac{\sum_{k \in \mathbb{I}_{ij}} c_k^2 ||\Psi_k||^2}{\sum_{k > 0} c_k^2 ||\Psi_k||^2}$$

 \mathbb{I}_{ij} is the set of bases with only x_i and x_j involved

Sensitivity information comes free with PC surrogate,

but not with piecewise PC

$$g(x_1,\ldots,x_d) = \sum_{k=0}^{K-1} c_k \Psi_k(\mathbf{x})$$

• Main effect sensitivity indices

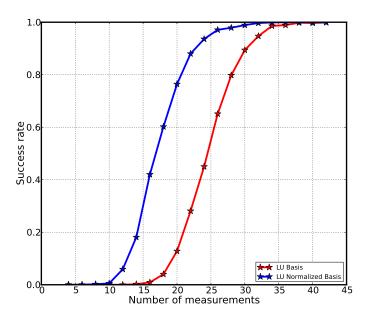
$$S_{i} = \frac{Var[\mathbb{E}(g(\mathbf{x}|x_{i}))]}{Var[g(\mathbf{x})]} = \frac{\sum_{k \in \mathbb{I}_{i}} c_{k}^{2} ||\Psi_{k}||^{2}}{\sum_{k > 0} c_{k}^{2} ||\Psi_{k}||^{2}}$$

Joint sensitivity indices

$$S_{ij} = \frac{Var[\mathbb{E}(g(\mathbf{x}|x_i, x_j))]}{Var[g(\mathbf{x})]} - S_i - S_j = \frac{\sum_{k \in \mathbb{I}_{ij}} c_k^2 ||\Psi_k||^2}{\sum_{k > 0} c_k^2 ||\Psi_k||^2}$$

 For piecewise PC, need to resort to Monte-Carlo estimation [Saltelli, 2002].

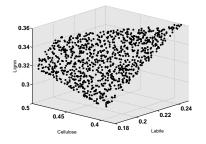
Basis normalization helps the success rate



Input correlations: Rosenblatt transformation

• Rosenblatt transformation maps any (not necessarily independent) set of random variables $\lambda = (\lambda_1, \dots, \lambda_d)$ to uniform i.i.d.'s $\{x_i\}_{i=1}^d$ [Rosenblatt, 1952].

$$\begin{aligned}
 x_1 &= F_1(\lambda_1) \\
 x_2 &= F_{2|1}(\lambda_2|\lambda_1) \\
 x_3 &= F_{3|2,1}(\lambda_3|\lambda_2,\lambda_1) \\
 \vdots \\
 x_d &= F_{d|d-1,...,1}(\lambda_d|\lambda_{d-1},...,\lambda_1)
 \end{aligned}$$



• Inverse Rosenblatt transformation $\lambda = R^{-1}(x)$ ensures a well-defined input PC construction

$$\lambda_i = \sum_{k=0}^{K-1} \lambda_{ik} \Psi_k(\mathbf{x})$$

• Caveat: the conditional distributions are often hard to evaluate accurately.

Strong discontinuities/nonlinearities challenge global polynomial expansions

- Basis enrichment [Ghosh & Ghanem, 2005]
- Stochastic domain decomposition
 - Wiener-Haar expansions,
 Multiblock expansions,
 Multiwavelets, [Le Maître et al, 2004,2007]
 - also known as Multielement PC [Wan & Karniadakis, 2009]
- Smart splitting, discontinuity detection
 [Archibald et al, 2009; Chantrasmi, 2011; Sargsyan et al, 2011; Jakeman et al, 2012]
- Data domain decomposition,
 - Mixture PC expansions [Sargsyan et al, 2010]
- Data clustering, classification,
 - Piecewise PC expansions

Piecewise PC expansion with classification

- Cluster the training dataset into non-overlapping subsets \mathcal{D}_1 and \mathcal{D}_2 , where the behavior of function is smoother
- Construct global PC expansions $g_i(x) = \sum_k c_{ik} \Psi_k(x)$ using each dataset individually (i = 1, 2)
- Declare a surrogate

$$g_s(\mathbf{x}) = \begin{cases} g_1(\mathbf{x}) & \text{if } \mathbf{x} \in^* \mathcal{D}_1 \\ g_2(\mathbf{x}) & \text{if } \mathbf{x} \in^* \mathcal{D}_2 \end{cases}$$

* Requires a classification step to find out which cluster *x* belongs to. We applied Random Decision Forests (RDF).

Caveat: the sensitivity information is harder to obtain.